A short introduction to INLA* and R-INLA * Integrated Nested Laplace Approximation

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Workshop: Theory and practice of INLA and SPDE

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Plan for this talk

• understand which models are suitable for estimation with INLA

• understand the basic mechanism of Integrated Nested Laplace Approximation

· learn how to run some simple models in R-INLA and how to conduct prediction

 \Rightarrow general theoretical concepts, and practical implementation in R-INLA

Some important resources for R-INLA with SPDE

http://www.r-inla.org/

- newsfeed
- discussion forum
- FAQ, examples and tutorials

A great book soon to come, already available for free online :

Advanced Spatial Modeling with Stochastic Partial Differential Equations Using R and INLA by E.T. Krainski, V. Gómez-Rubio, H. Bakka, A. Lenzi, D. Castro-Camilo, D. Simpson, F. Lindgren and H. Rue https://becarioprecario.bitbucket.io/spde-gitbook/

1 What models can we handle with INLA?

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8 R-INLA in practice Prediction with R-INLA Several practical examples

What kind of data/models can we fit with INLA?

- generalized additive (mixed) regression models
 - \Rightarrow explain observed response variable through covariates and random effects :

$$y \sim \underline{\beta_0 + \beta_1 \times \mathrm{covar.1} + \ldots + \beta_m \times \mathrm{covar.m}}$$
 + random effects

fixed effects

linear predictor

- large variety of response distributions for y
 - (log-)gaussian : (log) y =linear predictor+Gaussian error
 - Poisson : log(intensity) = log(𝔅y) = linear predictor
 - generalized Pareto : log(α-quantile of y)= linear predictor

and many more (gamma, skew normal, t, binomial), see inla.list.models("likelihood")

• random effects ⇒ high flexibility :

- model nonlinear trends with respect to space/time/covariates
- model longitudinal effects, space/time dependence
- structured (e.g., space/time dependence) or unstructured (e.g., measurement error)

 \Rightarrow capture variability in response not explained by fixed effects and response distribution

 \Rightarrow SPDE approach very useful for continuous random effects (Haakon's tutorial) see inla.list.models("latent")

Bayesian inference

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INLA uses a Bayesian framework :

 we put Gaussian prior distributions on fixed effects β_j and random effects, which can be (relatively) uninformative or incorporate expert knowledge

 \Rightarrow linear predictor is multivariate Gaussian

 \Rightarrow when we have only few data, prior has strong influence

- we can estimate some hyperparameters controling
 - signal-to-noise ratio, smoothing \Rightarrow precision parameters
 - range of dependence over space/time/..., e.g. range of Matérn correlation
 - · shape of the response distribution, e.g. gamma shape, skew-normal

see inla.list.models("prior")

Bayesian hierarchical formulation of models

$$egin{aligned} oldsymbol{ heta} &\sim \pi(oldsymbol{ heta}) \ \mathbf{x} \mid oldsymbol{ heta} &\sim \mathcal{N}(oldsymbol{0}, oldsymbol{Q}(oldsymbol{ heta})^{-1}) \ \mathbf{y} \mid \mathbf{x}, oldsymbol{ heta} &\sim \prod_i \pi(y_i \mid \eta_i(\mathbf{x}), oldsymbol{ heta}) \end{aligned}$$

hyperparameters latent Gaussian components likelihood of observations

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- univariate likelihood $\pi(y_i \mid \eta_i(\mathbf{x}), \boldsymbol{\theta})$ for data
- observation/projection matrix A links latent Gaussian components to observations :

$$\eta(\mathbf{x}) = (\eta_1(\mathbf{x}), \dots, \eta_n(\mathbf{x})) = \mathbf{A}\mathbf{x}$$
 where $y_i \sim \eta_i(\mathbf{x})$

- $Q(\theta)$ is called precision matrix (= inverse of variance-covariance matrix)
- R-INLA's speed, even in very high dimension, is based on
 - using sparse matrices for Q with mostly 0 entries
 - few non-0 entries in each row of A
 - conditional independence of data y given the linear predictor η and θ
 (⇒ sparsity is preserved during matrix computations in R-INLA)

Example : linear model with two covariates z_1 , z_2

 $\begin{array}{l} y_i \mid (\boldsymbol{x}, \boldsymbol{\theta}) \sim \mathcal{N}(\eta_i(\boldsymbol{x}), \sigma^2), \quad i = 1, \dots, n \\ \text{linear predictor } \eta_i(\boldsymbol{x}) = \beta_0 + \beta_1 z_{1i} + \beta_2 z_{2i} \\ \text{latent Gaussian vector } \boldsymbol{x} = (\beta_0, \beta_1, \beta_2) \end{array}$

• Gaussian likelihood
$$\pi(y \mid \eta_i(\mathbf{x}), \theta) = \frac{1}{\sqrt{2\pi \exp(-\theta)}} \exp\left(\frac{1}{2} \frac{y - \eta(\mathbf{x})}{\exp(-\theta)}\right)$$

•
$$\mathbf{x} \sim \mathcal{N}(\mathbf{0}, \mathbf{Q}^{-1})$$
 with $\mathbf{Q} = \begin{pmatrix} \tau_0 & 0 & 0 \\ 0 & \tau_\beta & 0 \\ 0 & 0 & \tau_\beta \end{pmatrix}$

for fixed effects, we usually fix the precision hyperparameters τ_0 , τ_β to a very low value, e.g. $10^{-3}~(\Rightarrow$ non-informative priors)

• hyperparameter $\theta = \log(1/\sigma^2) \in \mathbb{R}$ (log-precision of Gaussian likelihood), for instance with an exponential prior with rate λ on σ such that

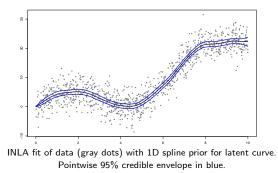
$$\pi(heta) = rac{\lambda}{2} \exp\left(-\lambda \exp(- heta/2) - heta/2
ight), \quad \lambda > 0$$

(Penalized Complexity Prior, see later talks)

Example : estimating a nonlinear regression curve

1D SPDE models are random effects in R-INLA, useful to estimate continuous curves :

- spline-like model with different degrees of smoothness
- Bayesian framework allows using many knots and correlating spline coefficients using a Matérn-like correlation model
- two hyperparameters : range, precision



What random effect models are available in R-INLA?

```
Output of inla.list.models("latent") :
```

ar	Auto-regressive model of order $p(AR(p))$
ar1	Auto-regressive model of order 1 (AR(1)
besag	The Besag area model (CAR-model)
generic	A generic model
iid	vector of independent effects
meb	Berkson measurement error model
ou	The Ornstein-Uhlenbeck process
rw1	Random walk of order 1
rw2	Random walk of order 2
rw2d	Thin-plate spline model
slm	Spatial lag model
spde2	A SPDE2 model

and many more ...

Instances of such models can be linked together through a group model (AR, RW, exchangeable, ...).

Example : first-order random walks

First-order Gaussian random walks (rw1 in R-INLA) are useful prior models for capturing nonlinear covariate effects or time trends.

They can be interpreted as very special and simple cases of 1D SPDE models.

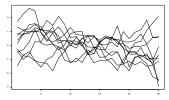
A rw1 is defined intrinsically over *m* classes through its "innovations"

$$x_{i+1} - x_i \sim \mathcal{N}(0, 1/\tau_{RW}), \quad i = 2, \dots, m.$$

It has one hyperparameter τ_{RW} measuring precision.

To make it identifiable, we have to impose a constraint such as $\sum_{i=1}^{m} x_i = 0$.

For easier interpretation, we may rescale to marginal distribution $\mathcal{N}(1/\tau_{RW})$ (i.e., control precision τ_{RW} averaged over all x_i).



10 realizations of a constrained rw1 over $t = 1, \ldots, 25$, rescaled, $1/\tau_{RW} = 1$

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Estimating Bayesian additive regression models

We obtain the posterior estimations through Bayes' formula.

▲ Usually, we cannot calculate the posterior estimations in closed form :

- posterior densities $\pi(\theta_j \mid \mathbf{y}), \pi(x_k \mid \mathbf{y}), \pi(\eta_i \mid \mathbf{y})$
- posterior mean estimates $\mathbb{E}(\theta_j \mid \mathbf{y})$, $\mathbb{E}(x_k \mid \mathbf{y})$, $\mathbb{E}(h(\eta_i) \mid \mathbf{y})$

 \bigwedge Moreover, the latent Gaussian components x are often very high-dimensional.

- \Rightarrow use numerical approximation of complicated integrals :
 - Markov-chain Monte-Carlo : iteratively simulate and update values x^(j), θ^(j)
 ⇒ generate a large representative sample of posterior distribution
 - theoretical convergence garantueed, but often too slow and unstable in practice \bigwedge mixing (i.e., exploration of the space of parameters) can be too slow
 - Integrated Nested Laplace Approximation : use astute numerical integration

Laplace approximation

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Suppose that we want to calculate

$$\int_{-\infty}^{\infty} \exp(g(\boldsymbol{x})) \, \mathrm{d}\boldsymbol{x}$$

where x is a vector with many components, such as the latent Gaussian components.

If g has maximum at x_0 and its values decrease fast and smoothly around x_0 , we can replace g(x) by a curvature approximation around x_0 :

$$g(\mathbf{x}) \approx g(\mathbf{x}_0) + 0.5(\mathbf{x} - \mathbf{x}_0)' \mathbf{H}(g)(\mathbf{x}_0)(\mathbf{x} - \mathbf{x}_0)$$

with Hessian matrix H containing second partial derivatives of g at x_0 .

Using this approximation, the function to integrate becomes an (unnormalized) Gaussian density :

$$\int_{-\infty}^{\infty} \exp(g(\mathbf{x})) d\mathbf{x} \approx (2\pi)^{d/2} |\mathbf{H}(g)(\mathbf{x}_0)|^{-1/2} \exp(g(\mathbf{x}_0))$$

In practice, it remains to determine x_0 and $H(g)(x_0)$ \Rightarrow iterative Newton-Raphson algorithm.

A schematic overview of how INLA works

All hyperparameters are fixed : skip 1 and 2.2. Exactly 1 hyperparameter : skip 1.2.

() Calculate posterior density of hyperparameter θ_j :

$$\pi(\theta_j \mid \mathbf{y}) = \int \int \pi(\boldsymbol{\theta}, \mathbf{x} \mid \mathbf{y}) \, \mathrm{d}\mathbf{x} \mathrm{d}\boldsymbol{\theta}_{-j}$$

- 1 dx: use Laplace approximation for each configuration θ
- **2** $d\theta_{-j}$: use numerical integration

2 Calculate posterior density of latent components x_i or predictors η_i :

$$\pi(\mathbf{x}_i \mid \mathbf{y}) = \int \int \pi(\mathbf{x}, \mathbf{\theta} \mid \mathbf{y}) \mathrm{d}\mathbf{x}_{-i} \, \mathrm{d}\mathbf{\theta}$$

- **1** $d\mathbf{x}_{-i}$: use Laplace approximation (strategy="laplace"), but often too expensive
 - cheap but often less accurate : use conditional Gaussian from dx above (strategy="gaussian")

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- R-INLA default is a simplified Laplace approximation (strategy="simplified.laplace")
- 2 $d\theta$: use numerical integration

R-INLA proposes three variants for numerical integration in 1.2 and 2.2 :

- grid of θ configurations (int.strategy="grid", most costly)
- central composite design (int.strategy="ccd", the default)
- only the mode of $\pi(\theta \mid \mathbf{y})$ (empirical Bayes, int.strategy="eb")

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Prediction with R-INLA

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The output of inla(...) contains the posterior estimates of fitted values for observations, but we often want to predict where no observations are available :

- spatial interpolation
- temporal prediction
- fill missing data values

Another simple possibility :

 "construct" the posterior mean of the linear predictor for new unobserved data with observation matrix A_{new} by transforming the posterior mean of latent Gaussian components,

 $\hat{\eta}_{\mathsf{new}} = A_{\mathsf{new}} \, \mathbb{E}(\mathbf{x} \mid \mathbf{y})$

2 transform to the scale of the response if necessary

 \bigwedge propagation of posterior uncertainty of **x** given **y** in Step 2 is disregarded ! e.g. with log-link, exp($\hat{\eta}_{new}$) will tend to underestimate the response values

Yet another possibility : generate a large sample from the posterior distribution of model components using inla.posterior.sample(...), and then conduct Monte-Carlo calculation of predictions etc.

Practical examples

We will now treat several code examples \Rightarrow open <code>examples.R</code> in <code>RStudio</code>.

Example 1 : monthly temperatures in Nottingham 1920–1939

nottem dataset of R

Objective : model seasonal behavior

• use monthly random walk (m = 12) to capture seasonality in data $\Rightarrow 12$ latent Gaussian variables x_1, \ldots, x_{12} with

$$x_{t+1} - x_t \sim \mathcal{N}(0, 1/\tau_{RW})$$

 random walk values for January and December should be linked ⇒ make the random walk cyclic by setting

$$x_1 - x_{12} \sim \mathcal{N}(0, 1/\tau_{RW})$$

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Example 2 : Daily air quality data for New York, May-September 1973

airquality dataset of $R\ m$ containing Ozone (with missing values), temperature, wind speed,...

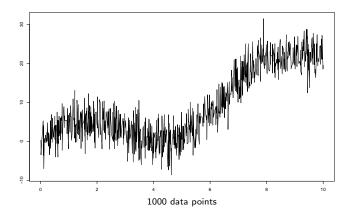
Objective : predict missing values of Ozone

- model 1 : temperature and wind as linear covariates (fixed effects)
- model 2 : temperature and wind as nonlinear covariates
 ⇒ use rw1 over classes of covariate values
- model 3 : temperature and wind (fixed effects), ar1 for time dependence
- could combine two of the preceding models, e.g. model 2 and model 3
- \Rightarrow select best model through cross-validation of root-mean squared error (or of any other good metric of predictive performance)

Example 3 : 1D SPDE for nonlinear regression curve

simulated dataset

Objective : fit a smooth nonlinear curve



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